

LECTURE 9

TODAY :

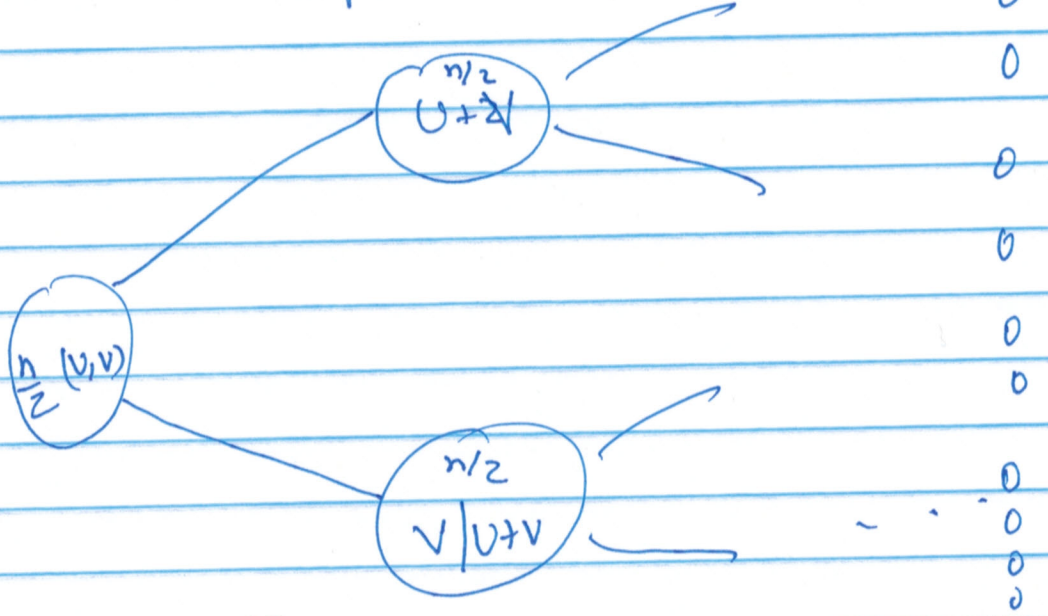
POLAR CODING (contd.)

- DECODING
- TOWARDS ANALYSIS (of Entropy Theorem)
 - MARTINGALES
 - LOCAL VS. GLOBAL POLARIZATION



Review of last lecture

1) Want to compress i.i.d Bern(p) variable



2) Theorem: $H(p) \leq \sum A_i B_i$ poly n

Key
 (2) Theorem (to be proved):

$$\forall p, c < \infty, \exists \beta < 1 \text{ s.t. } \forall t$$

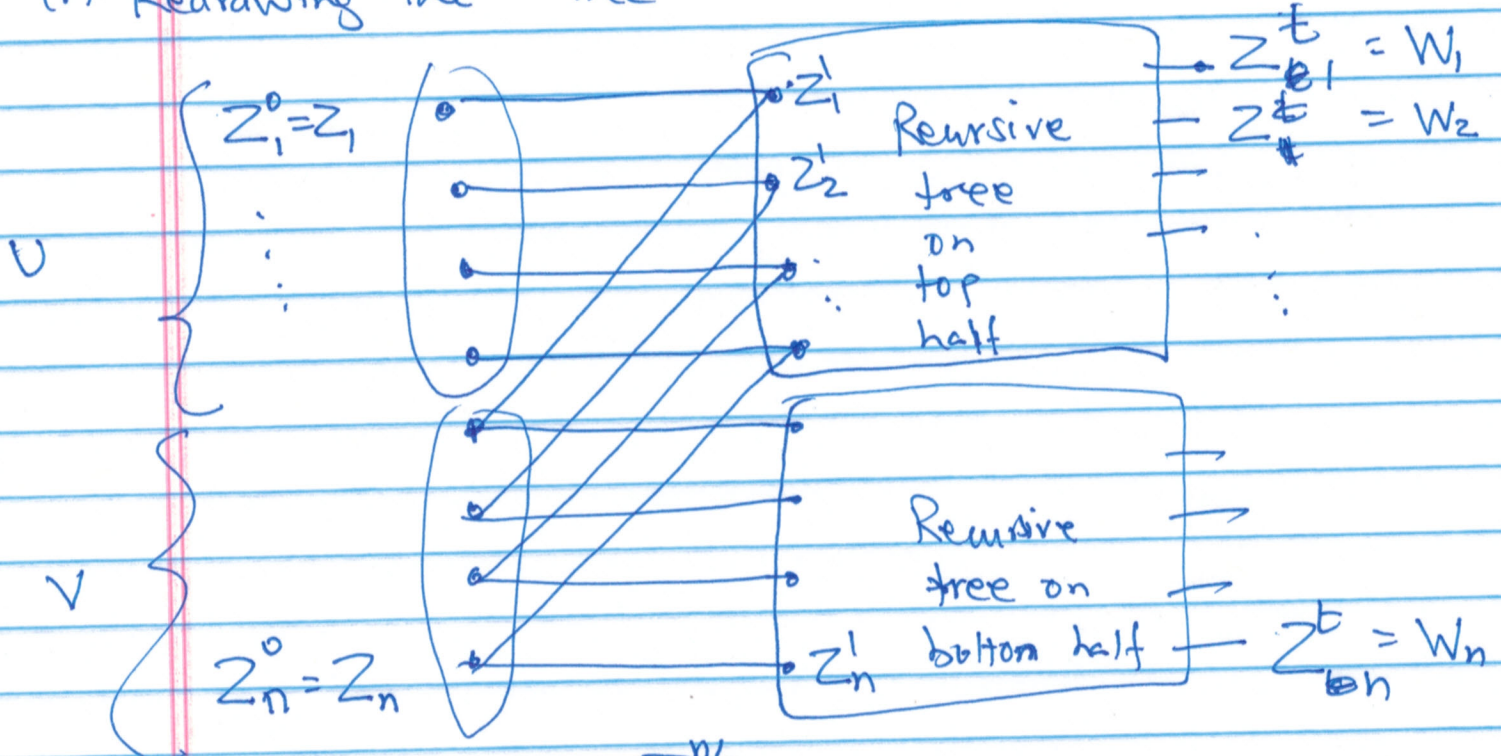
$$\Pr_{j \in [2^t]} \left[H(w_j | W_{2^t}) \in (c^{-t}, 1 - c^{-t}) \right] \leq O(\beta^t)$$

(3) How to decode?

$$S \triangleq \{j \mid H(w_j | W_{2^t}) \geq c^{-t}\}$$

DECODING

(i) Redrawing the "tree"

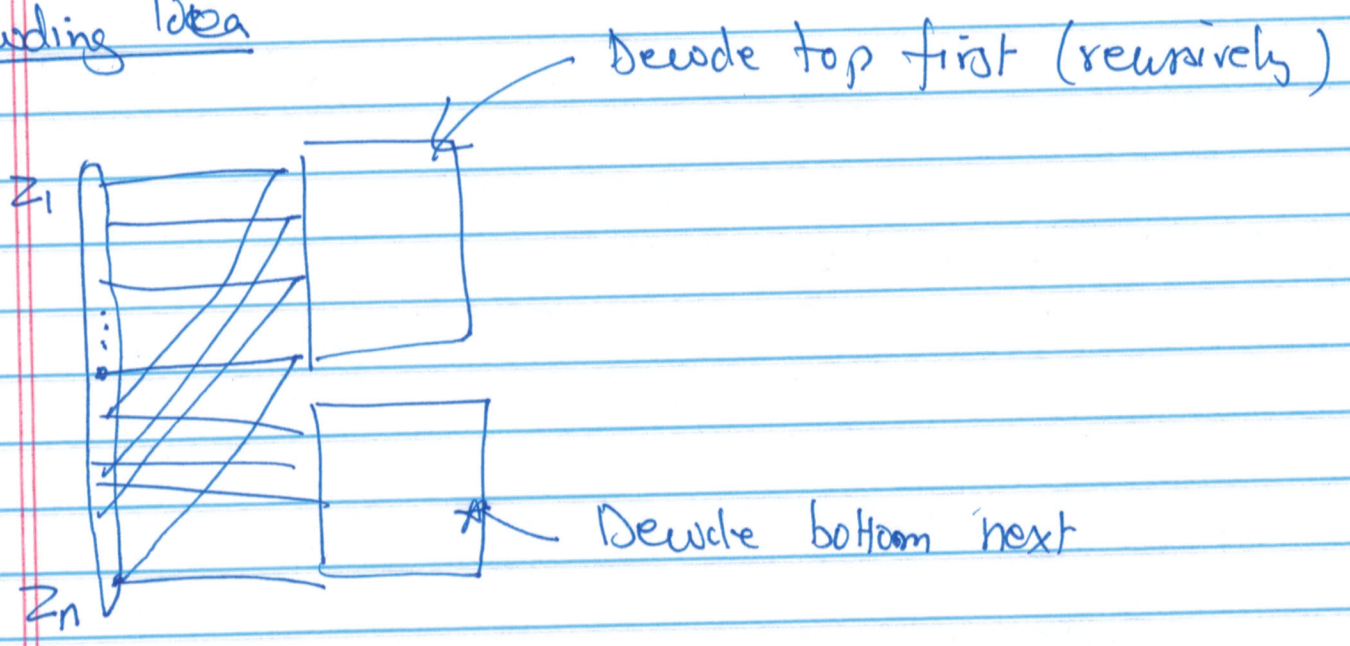


$$z = (u, v) \quad u, v \in \mathbb{F}_2^{n/2}$$

$$W = P_n(u, v) = \left(P_{n/2}(u+v), P_{n/2}(v) \right)$$

Output W_s .

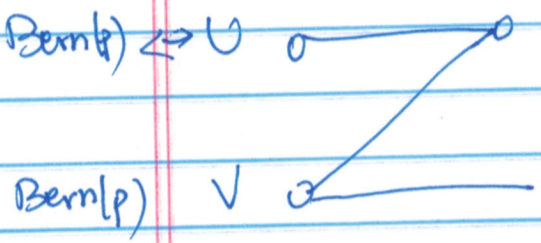
Decoding Idea



top decoding: "obvious".

- bottom?: Need to exploit knowledge of top decoding!
(Else impossible!)

- how? can compute revised "bias" of bits. why?



$$\Pr[V=1 | U+V=1] = \frac{1}{2}$$

$$\Pr[V=1 | U+V=0] \approx O(p^2) \text{ [as } p \rightarrow 0 \text{]}$$

$$= \frac{p^2}{p^2 + (1-p)^2}$$

- But now w bits not identical!:

$$\Pr[Z_{\frac{n}{2}+1} = 1] = \frac{1}{2}, \Pr[Z_{\frac{n}{2}+2} = 1] = p^2 \text{ et cetera}$$

DECODE (W; (P₁, ..., P_n))

$W \in \{0, 1, ?\}^n ; 0 \leq P_i \leq 1$

- $\hat{U} \hat{V} = \text{Decode} (W[1..n/2]; q_1 \dots q_{n/2})$

$q_i \triangleq P_i(1 - P_{n/2+i}) + P_{n/2+i}(1 - P_i)$

- $\hat{V} = \text{Decode} (W[n/2+1..n]; r_1 \dots r_{n/2})$

$r_i = \frac{P_{n/2+i}(1 - P_i)}{q_i}$ if $(\hat{U} + \hat{V})_i = 1$

$= \frac{P_i P_{n/2+i}}{1 - q_i}$ if $(\hat{U} + \hat{V})_i = 0$

- Output (\hat{U}, \hat{V})

Base Case?

Decode (W; P_i) n=1

if $W_i = 0/1$ output W_i ,

else ($W_i = ?$) output 1 if $P_i \geq 1/2$

0 o.w.



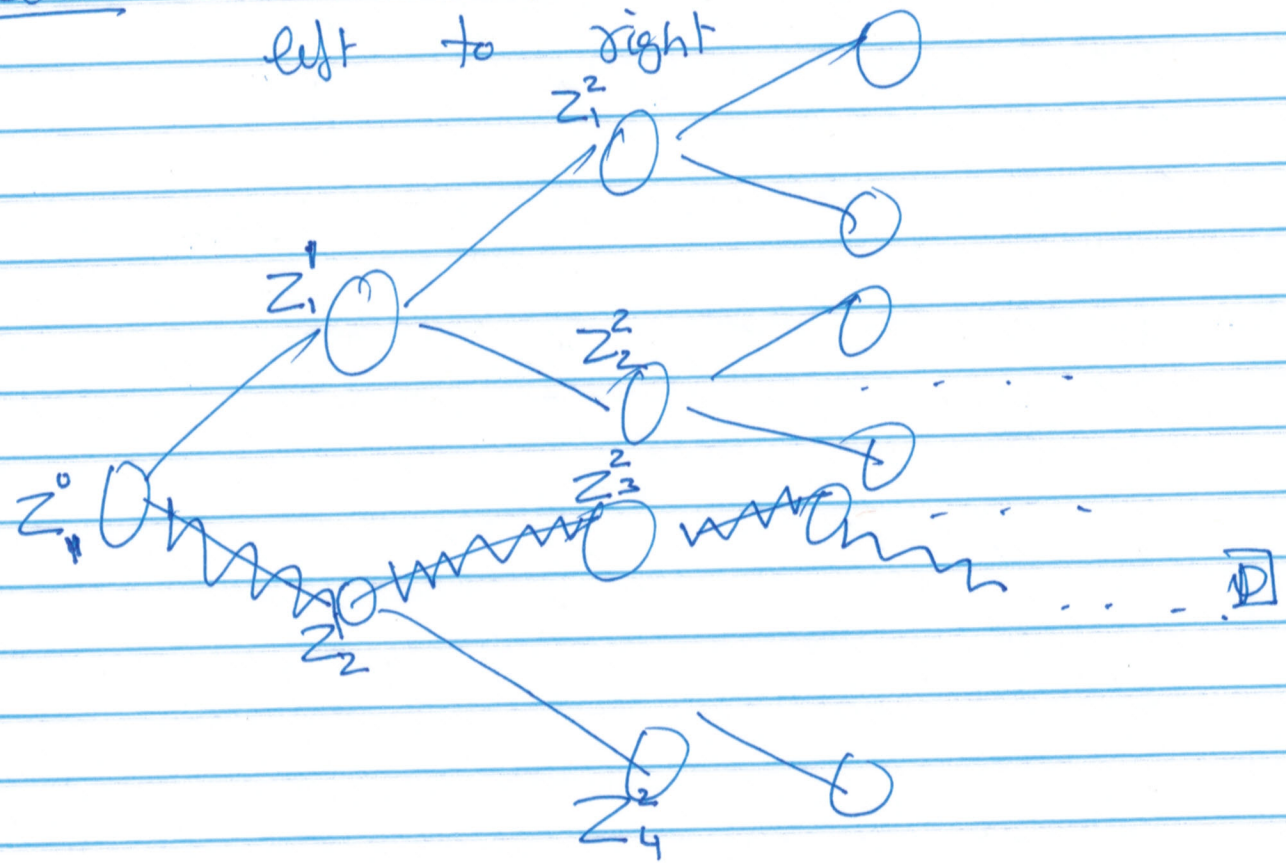
Key Analysis Ingredient:

- Biases are correct conditioned on things above, as we move left to right.



Proof of Polarization Theorem

"Perspective": take random walk on tree from left to right



- follow entropy $H(z_j^i | z_{z_j^i}^i) \triangleq X_i$

- Theorem (to be proved)

$$Pr_{\text{walk}} [X_i \in (c^{-t}, 1-c^{-t})] \leq O(\beta^t)$$

$\forall \alpha$
Fact: $\mathbb{E}[X_i | X_0 \dots X_{i-1} = \alpha] = \alpha$ "Martingale"
 $(\leftarrow (u, v) \rightarrow (u+v, v) \text{ is invertible})$.

Our martingale is "bounded": $X_i \in [0, 1] \forall i$.

What we see (almost imm.)

Local Polarization:

① $\forall \tau > 0 \exists \sigma > 0$ s.t. [Variance in middle]

$\text{Var}(X_i | X_1 \dots X_{i-1}) \geq \sigma$ if $X_{i-1} \in (\tau, 1-\tau)$.

~~$0 < \sigma < \tau$~~

② $\forall c < \infty \exists \tau > 0$ s.t. $\forall i$ [Suction at ends]

$\Pr[X_i > \frac{X_{i-1}}{c}] \geq \tau$ if $X_{i-1} \leq \tau$

[similarly for $\bar{X}_i = 1 - X_i$].

Key Lemma: Local Polarization \Rightarrow Strong Polarization
(condition of Theorem).

Lemma 8: Our Martingale is Locally Polarizing

Proof Idea: Pretend it's all prob. (not conditional)

① if $U, V \sim \text{Bern}(p)$

$\text{bias}(U+V) = 2p(1-p) > p \Rightarrow \text{Variance}$.

(2) (Suction at high end)

$$p = \frac{1}{2} - \delta \quad \delta \rightarrow 0$$

$$\Rightarrow H(p) = 1 - \delta^2$$

$$\& p' = 2p(1-p) = \frac{1}{2} - \delta^2$$

$$\Rightarrow H(p') = 1 - \delta^4$$

$$\cdot (\bar{X}_i = \delta^2 \Rightarrow \bar{X}_{i+1} = \delta^4 \text{ more than good enough})$$

Suction at low end:

$$p = \delta \Rightarrow H(p) \approx \delta \cdot \log \frac{1}{\delta}$$

$$p' = 2p(1-p) \approx 2p$$

$$H(p') \approx 2\delta \log \frac{1}{2\delta}$$

$$\Rightarrow 2H(p) - H(p') \approx 2\delta \log \frac{1}{\delta} - 2\delta \log \frac{1}{2\delta}$$

$$\approx 2\delta \cdot \frac{H(p)}{\log H(p)} = o(H(p))$$



Proof of Key Lemma

Part 1

$$\phi_t \triangleq \min \left\{ \sqrt{X_t}, \sqrt{1-X_t} \right\}$$

then $\mathbb{E}[\phi_{t+1} | \mathcal{F}_t] \leq \beta' \cdot \phi_t$

• Proof of claim: Calculation

$$\Rightarrow \mathbb{E}[\phi_t] \leq (\beta')^t \Rightarrow \Pr[\phi_t \geq (\beta')^{t/2}] \leq (\beta')^{t/2}$$

$$\Rightarrow \Pr[X_t \in ((\beta')^{t/2}, 1 - (\beta')^{t/2})] \leq (\beta')^{t/2}$$

not good enough

good!

② if $X_t \leq (\beta')^{t/2}$ then $X_{2t} < C^{2t}$ whp.

roughly: while $\exists T_0$ s.t. with

$$X_{t'} > T_0 \quad \Pr[X_{t'+1} < \frac{X_{t'}}{C^{10}}] \gg \frac{1}{2}$$

$$\Pr[\exists t' \in [t, 2t] \text{ s.t. } X_{t'} > T_0]$$

$$\leq \frac{(\beta')^{t/2}}{T_0} \quad [\text{Doob's Ineq.}]$$

if $X_{t'} < T_0 \forall t'$... then falls w.p.

$\frac{1}{2}$ by C^{10} ; doubles w.p. $\frac{1}{2}$