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# Infinite Multiway Mixture with Factorized Latent Parameters

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## Abstract

In this paper, we develop an infinite multiway mixture model, whose parameters are represented as a tensor factorization. We define a D-way Poisson mixture, where a large observed tensor  $X$  is generated by the mixture proportions  $\pi_d$  and a smaller latent tensor  $\Theta$ , which is represented as a factorization of  $M$  latent factors  $\Theta_m$  of varying dimensionalities. We first derive an EM algorithm for the finite mixture. Then, we formulate an infinite multiway mixture, and propose an MCMC method to sample the assignments.

## 1 Introduction

Clustering has been a primary problem in Bayesian nonparametrics that led to the development of a literature on DP mixtures. Multiway clustering is a problem that has recently gained attention. In [1], the multiway clustering problem is formulated with reference to earlier work on using hypergraphs to approach VLSI and PCB clustering placement problem. As elaborated in [2], a hypergraph is a general representation that contains hyperedges of any size that relate any combination of entities. In [3], a general multiway framework is presented to handle various kinds of hyperedges. Our model assigns D-tuples of objects to D-tuples of clusters, thus only involves D-way hyperedges that relate D objects of different types.

We use D-way tensors to represent the variables. In [4], a probabilistic tensor factorization framework is presented for multiway analysis. We use a similar framework to represent the latent D-way tensor of component parameters in the multiway mixture. In [5], an MCMC method was proposed for nonparametric biclustering problem. Biclustering is a special case of D-way clustering for  $D = 2$ , thus its solution can be applied to the general multiway problem.

In the following sections, we first formulate the multiway clustering problem as a finite mixture, and present a variational inference method. Then, we present an MCMC inference method to use with the infinite mixture model.

## 2 The D-way Poisson mixture model

In our problem, we have D types of objects, and  $N_d$  objects from each object type  $d \in \{1, \dots, D\}$ . Each of the observations is given for a D-tuple of these objects, together making up the D-way tensor  $X$  with sizes  $N_1, \dots, N_D$  in its D dimensions. We model  $X$  as a Poisson mixture of latent parameters in a smaller D-way tensor  $\Theta$  with sizes  $K_1, \dots, K_D$ . Here,  $K_d$  is the number of clusters for a given dimension, and in general it is significantly smaller than  $N_d$ .

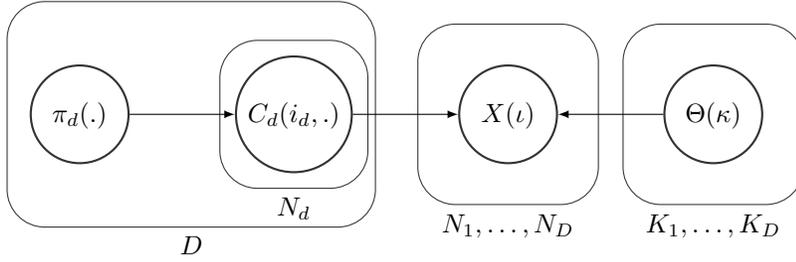


Figure 1: The D-way mixture model

A D-way tensor is indexed by an ‘index set’ of D indices. Each observation in the tensor  $X$  is denoted  $X(\iota)$  where  $\iota$  is the index set  $\{i_1, \dots, i_D\}$  and  $i_d \in \{1, \dots, N_d\}$ . Similarly, the tensor  $\Theta$  is indexed by  $\kappa$  that is the index set  $\{k_1, \dots, k_D\}$  where  $k_d \in \{1, \dots, K_d\}$ .

## 2.1 Layer assignments

In D-way clustering, for each object type d,  $N_d$  objects of this type are to be assigned to the corresponding  $K_d$  clusters. For  $D = 1$ , single observations in  $X$  are assigned to single parameters in  $\Theta$ . For  $D = 2$ , rows of observations in  $X$  are assigned to rows of parameters in  $\Theta$ , and columns to columns. When  $D = 3$ , matrices in three different orientations from  $X$  are assigned to matrices of corresponding orientations in  $\Theta$ . For the general case, (D-1)-way tensors in  $X$  are assigned to (D-1)-way tensors in  $\Theta$ . We call such a (D-1)-way tensor a ‘layer’, and indicate it by a dimension d, and a value for its index  $j_d$ . The layer’s orientation is given by d, and its placement inside the tensor by  $j_d$ .

We call  $C_d$  an indicator variable. For each orientation d, that  $C_d(i_d, k_d) = 1$  indicates that the layer at  $i_d$  of  $X$  is assigned to the layer at  $k_d$  of  $\Theta$ . A single observation  $X(\iota)$  is thus assigned by the indicators  $C_1, \dots, C_D$  to the layers at the indices  $k_1, \dots, k_D$  of  $\Theta$ . These indices form the set  $\kappa$ , and as a result, the observation is assigned to the latent parameter at  $\Theta(\kappa)$ .

In the mixture model, we assume that each observation  $X(\iota)$  is Poisson distributed with the intensity as the latent parameter  $\Theta(\kappa)$  to which it is assigned.

$$X(\iota) \mid \Theta, C \sim \prod_{\kappa} \mathcal{PO}(\Theta(\kappa))^{\prod_{d=1}^D C_d(i_d, k_d)}$$

We model each of the vectors  $C_d(i_d, \cdot)$  by a discrete distribution  $\pi_d(\cdot)$  of size  $K_d$ . This vector is in turn modeled by a symmetric Dirichlet prior with concentration  $\alpha_d$ . The full model is shown in Figure 1.

$$\begin{aligned} C_d(i_d, \cdot) \mid \pi_d &\sim \text{Discrete}(\pi_d(\cdot)) \\ \pi_d(\cdot) &\sim \text{Dir}\left(\frac{\alpha_d}{K_d}, \dots, \frac{\alpha_d}{K_d}\right) \end{aligned}$$

## 2.2 Representing the latent tensor

Up to this point, we have described a general D-way mixture model. What makes our model specific is the representation of the latent tensor  $\Theta$ . We assume that  $\Theta$  is a function of other M latent factors  $\Theta_m$  of different dimensionalities.

$$\Theta(\kappa) = \sum_{\beta} \prod_{m=1}^M \Theta_m(\gamma_m)$$

The factorization is summed over a set of indices  $\beta$  to get the latent tensor  $\Theta$  indexed by  $\kappa$ . Here,  $\beta$  denotes an ‘additional’ set of indices  $\{k_{D+1}, \dots, k_{D+\Delta}\}$  that extends the ‘original’ set denoted by  $\kappa$ . The union of these two gives the full set of indices  $\kappa \dot{\cup} \beta = \{k_1, \dots, k_{D+\Delta}\}$ . Each of the M factors is indexed by  $\gamma_m$ , such that  $\cup_{m=1}^M \gamma_m = \kappa \dot{\cup} \beta$ .

We consider the factors  $\Theta_m$  as the actual hidden parameters, and put on them a Gamma prior, which is conjugate with the Poisson distribution.

$$\Theta_m(\gamma_m) \sim \mathcal{G}(A, \frac{B}{A})$$

### 2.3 Partitioning a factor's indices

The index set  $\gamma_m$  of a factor is partitioned into three disjoint sets in two consecutive steps as follows:

$$\gamma_m = \eta_m \dot{\cup} \beta_m = \eta_m \dot{\cup} \sigma_m \dot{\cup} \lambda_m$$

In the first step, it is partitioned into its ‘original’ indices  $\eta_m = \gamma_m \cap \kappa$  and ‘additional’ indices  $\beta_m = \gamma_m \cap \beta$ . We then introduce  $\beta_{-m} = \cup_{m' \neq m} \beta_{m'}$  to denote the additional indices that belong to any factor other than  $m$ . In the second step,  $\beta_m$  is further partitioned into the indices shared with other factors  $\sigma_m = \beta_m \cap \beta_{-m}$  and the indices that are not shared  $\lambda_m = \beta_m \setminus \beta_{-m}$ .

### 2.4 Determining the parameters of the model

For a given  $D$ , a finite multiway mixture model is selected by determining the following:

1. The sizes  $\{K_1, \dots, K_D\}$  of the latent tensor  $\Theta$  (the number of clusters for all object types).
2. The concentration parameters  $\alpha_1, \dots, \alpha_D$  for each of the object types.
3. The number of factors  $M$ , and the index sets  $\gamma_1, \dots, \gamma_M$  for each of these factors. When these are given, we can also determine the set of additional indices  $\beta = \cup_{m=1}^M \gamma_m \setminus \kappa$ .
4. The prior parameters  $A$  and  $B$  for the factors  $\Theta_m$ .

Various factorizations of  $\Theta$  lead to different models. To mention two basic examples: When  $M = 1$  and  $\gamma_1 = \kappa$ , the latent tensor  $\Theta(\kappa)$  is modelled directly. When  $M = D$  and  $\gamma_d = \{k_d\}$ , the tensor  $\Theta$  is the product of  $D$  vectors  $\Theta_d(k_d)$ .

## 3 Variational inference for the finite mixture

We develop an Expectation-Maximization algorithm that involves the following steps

1. Calculate the expectation of  $p(C | X, \Theta, \pi)$ .
2. For each  $d \in \{1, \dots, D\}$ , calculate the  $\pi_d$  that maximizes  $p(X, C, \Theta, \pi)$ .
3. For each  $m \in \{1, \dots, M\}$ , calculate the  $\Theta_m$  that maximizes  $p(X, C, \Theta, \pi)$ .

In the expectation step (1), we use the posterior of the layer assignments.

$$p(C | X, \Theta, \pi) \propto \left\{ \prod_{\iota} \prod_{\kappa} \mathcal{PO}(X(\iota) | \sum_{\beta} \prod_{m=1}^M \Theta_m(\gamma_m))^{\prod_d C_d(i_d, k_d)} \right\} \left\{ \prod_d \prod_{i_d} \prod_{k_d} \pi_d(k_d)^{C_d(i_d, k_d)} \right\}$$

In the maximization step (2), we update  $\pi_d(k_d)$  by the equation:

$$\pi_d^*(k_d) = \frac{\frac{\alpha_d}{K_d} - 1 + \sum_{i_d} \mathbb{E}[C_d(i_d, k_d)]}{\alpha_d - K_d + N_d}$$

In the next step (3), we update  $\Theta_m(\gamma_m)$  by the following formula:

$$\Theta_m^*(\gamma_m) = \frac{A - 1 + \sum_{\iota} \{X(\iota) \frac{\prod_{d: k_d \in \lambda_m} K_d}{\sum_{\lambda_{m'} \neq \lambda_m} \Theta_m(\eta_m \cup \sigma_m \cup \lambda_{m'})}\} \mathbb{E}[\prod_d C_d(i_d, k_d)]}{\frac{B}{A} + \sum_{\iota} \{\sum_{\beta} \prod_{m' \neq m} \Theta_{m'}(\gamma_{m'})\} \mathbb{E}[\prod_d C_d(i_d, k_d)]}$$

For any factor  $\Theta_m(\gamma_m)$  with no additional indices (for each  $m$  where  $\beta_m = \emptyset$ ), the fraction in the numerator of the formula reduces to 1. When there is no factorization ( $M = 1$ ), the coefficient in the summation in the denominator also reduces to 1.

## 4 The infinite mixture and MCMC inference

By taking a finite D-way mixture, and bringing  $K_d \rightarrow \infty$  for some or all of the object types, we can obtain an infinite multiway mixture. We are developing an MCMC method for inferring the layer assignments in such a nonparametric multiway mixture model.

A variety of MCMC methods for DPM are presented in [6] including Gibbs sampling, Metropolis-Hastings updates and auxiliary parameters to handle both conjugate and non-conjugate priors. In [5], such a method is developed for a nonparametric biclustering model, which can be obtained from our D-way model for  $D = 2$ . The method proposed is based on a property which we will express in our terms as follows.

When  $C_d$  are given, for each  $\kappa$ , there is a set or a ‘block’ of observations that are assigned to it:

$$\{X(\iota) : \kappa\} = \{X(\iota) : \prod_d C_d(i_d, k_d) = 1\}$$

Conditional to  $\Theta$ , these blocks of observations are independent and thereby their likelihoods:

$$p(X | \Theta, C) \sim \prod_{\kappa} p(\{X(\iota) : \kappa\} | \Theta(\kappa))$$

In case of a conjugate prior, we can also integrate out  $\Theta$  to get the following:

$$p(X | C) \sim \prod_{\kappa} \int p(\{X(\iota) : \kappa\} | \Theta(\kappa)) p(\Theta(\kappa)) d\Theta(\kappa) = \prod_{\kappa} p(\{X(\iota) : \kappa\})$$

Using this property, an MCMC algorithm similar to [5] can be developed for the infinite D-way mixture.

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